



Medina-Gomez, C., Kemp, J. P., Trajanoska, K., Luan, J., Chesi, A., Ahluwalia, T. S., Mook-Kanamori, D. O., Ham, A., Hartwig, F. P., Evans, D. S., Joro, R., Nedeljkovic, I., Zheng, H. F., Zhu, K., Atalay, M., Liu, C. T., Nethander, M., Broer, L., Porleifsson, G., ... Rivadeneira, F. (2018). Life-course genome-wide association study meta-analysis of total body BMD and assessment of age-specific effects. *American Journal of Human Genetics*, 102(1), 88-102.  
<https://doi.org/10.1016/j.ajhg.2017.12.005>

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Link to published version (if available):  
[10.1016/j.ajhg.2017.12.005](https://doi.org/10.1016/j.ajhg.2017.12.005)

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# Life-course Genome-Wide Association Study Meta-analysis of Total Body BMD and

## Assessment of Age-specific Effects

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## Abstract

Bone mineral density (BMD) assessed by DXA is used to evaluate bone health. In children, total body (TB) measurements are commonly used; in older individuals, BMD at the lumbar spine (LS) and femoral neck (FN) is used to diagnose osteoporosis. To date, genetic variants in more than 60 loci have been identified as associated with BMD. To investigate the genetic determinants of TB-BMD variation along the life course and test for age-specific effects, we performed a meta-analysis of 30 genome-wide association studies (GWAS) of TB-BMD including 66,628 individuals overall and divided across five age-strata each spanning 15 years. We identified variants associated with TB-BMD at 80 loci, of which 36 have not been previously identified; overall they explain approximately 10% of the TB-BMD variance when combining all age groups and influence the risk of fracture. Pathway and enrichment analysis of the association signals showed clustering within gene-sets implicated in the regulation of cell growth and SMAD proteins; overexpressed in the musculoskeletal system; and enrichment in enhancer and promoter regions. These findings reveal TB-BMD as a relevant trait for genetic studies of osteoporosis, enabling the identification of variants and pathways influencing different bone compartments. Only variants in *ESR1* and close proximity to *RANKL* showed a clear effect dependency on age. This most likely indicate that the majority of genetic variants identified influence BMD early in life and their effect can be captured throughout the life course.

## Introduction

Osteoporosis is a disease characterized by low bone mass and microarchitectural deterioration of bone tissue leading to increased risk of fracture<sup>1</sup>. It is diagnosed through the measurement

of bone mineral density (BMD) utilizing dual-energy X-ray absorptiometry (DXA), which is the single best predictor of fracture<sup>1</sup>.

Bone is a dynamic tissue constantly undergoing resorption and formation. Bone mass increases steadily during childhood and markedly during adolescent growth<sup>2</sup>. Peak bone mass is attained at approximately the third decade of life. Thereafter, until about 50 years of age, BMD remains fairly stable, by virtue of the coupling between bone formation and resorption (e.g., bone remodeling). Subsequently, bone resorption exceeds the rate of bone formation, resulting in a decrease in BMD, particularly in women after the onset of menopause<sup>3</sup>.

The International Society for Clinical Densitometry recommends performing DXA measurements at the lumbar spine, femoral neck and total hip to diagnose osteoporosis in postmenopausal women and men who are 50 years or older<sup>4</sup>. Consequently, studies of BMD determinants are frequently based on measurements at these skeletal sites. By contrast, for the assessment of bone health in children and adolescents, total body (excluding head) and lumbar spine are the preferred sites to minimize measurement artifacts resulting from changing areas in growing bones<sup>4</sup>. Nevertheless, in elderly individuals degenerative changes in the spine can give elevated BMD readings<sup>5</sup>. Moreover, total body DXA scans have been obtained in a number of adult research cohorts, primarily to assess body composition. Therefore, the total body BMD (TB-BMD) measurement is the most appropriate method for an unbiased assessment of BMD variation in the same skeletal site from childhood to old age.

To date, nearly 80 independent genetic variants have been shown to be robustly associated with variability in bone parameters<sup>6-18</sup>. Most of these markers have been identified in studies

comprising tens of thousands of adult and elderly individuals with DXA-derived BMD measurements, although a few of them have been associated with BMD specifically in studies of pediatric cohorts<sup>8</sup>. Furthermore, several of the associated variants display significant site-specific effects, possibly reflecting differences in bone composition across skeletal sites (e.g., cortical bone vs. trabecular bone) or differential response to mechanical loading<sup>8</sup>. Moreover, genetic studies on measures from peripheral quantitative computed tomography (pQCT) and bone quantitative ultrasound, which provide additional information regarding bone size, geometry and (micro) architecture identified genetic variants that may have specific effects on bone properties that are poorly captured by conventional DXA measurements<sup>9-10</sup>.

Given the complex physiological processes underlying age-related changes in BMD across the life course, it is possible that genetic studies in more refined age groups will reveal variants in unreported loci as well as age-specific genetic effects. Thus, the purpose of this study was to identify gene variants associated with TB-BMD across the life span and investigate possible differences of genetic effects across age periods.

## Methods

### *TB-BMD GWAS meta-analyses*

#### Study Populations

##### Subjects

This study comprised 30 epidemiological studies comprising ~66,628 individuals from populations across America, Europe, and Australia, with a variety of designs (**Supplemental Data; Table S1**) and participant characteristics (**Table S2**). In summary, most participants came from population-based cohorts of European ancestry (86%), two cohorts comprising African-American individuals (2%) and other four studies holding a fraction of individuals from admixed background (14%). All research aims and the specific measurements have been approved by the correspondent Medical Ethical Committee of each participating study. Written informed consent was provided by all subjects or their parents in the case of children.

##### BMD measurement

Total body BMD ( $\text{g}/\text{cm}^2$ ) was measured by DXA following standard manufacturer protocols. As recommended by the International Society for Clinical Densitometry total body less head (TBLH) was the measurement used in pediatric cohorts<sup>4</sup> (e.g., 0-15 years). Detailed information on the assessments performed by each study can be found in **Table S1**.

##### GWAS data and imputation

All individuals included in this study had genome-wide array data. Quality control of genotypes is summarized in **Table S1**. To enable meta-analysis, each study performed genotype

imputation using the cosmopolitan (all ethnicities combined) 1000 genomes phase 1 version 3 (March 2012) reference panel, yielding ~ 30,000,000 SNPs for analysis. Three studies used the combined 1000 genomes and the UK10K reference panels as presented in **Table S1**.

## **Association Analysis**

TB(LH)-BMD was corrected for age, weight, height and genomic principal components (derived from GWAS data), as well as any additional study-specific covariates (e.g. recruiting center), in a linear regression model. For studies with non-related individuals, residuals were computed separately by sex, whereas for family-based studies sex was included as a covariate in the model. Finally, residuals were inverse normal transformed. The analyses were performed in each study for the overall population as well as in subgroups of individuals by age-strata, defined by bins of 15 years (i.e., 0-15 years, 15-30 years, 30-45 years, 45-60 years, and 60 or more years). SNP association was tested for autosomal variants, in which the additive effect of each SNP on the normalized BMD-residuals was estimated via linear regression.

## **Quality control of TB-BMD association summary statistics**

A centralized quality-control procedure implemented in EasyQC<sup>19</sup> was applied to all study-specific files of association results to identify cohort-specific issues. We excluded variants if they had missing information (e.g., missing association P-value, beta estimate, alleles, allele frequency), or nonsensical values (e.g., absolute beta estimates or standard errors >10, association P-values >1 or <0; or imputation quality < 0; infinite beta estimates or standard errors); minor allele frequency (MAF) less than 0.5%; imputation quality scores <0.4 (Impute2) or <0.3 (Minimac). Moreover, variants were flagged if they had large allele frequency



deviations from reference populations ( $>0.6$  for admixed studies and  $>0.3$  for ancestry-homogeneous studies).

### GWAS meta-analyses

In the first instance, no exclusion criteria based on ancestry were applied for the meta-analysis ( $N=66,628$ ). In addition, meta-analyses were carried out across age strata (minimum sample size per bin  $N=200$  for each study) comprising: 1) 0-15 years ( $N=11,807$ ), 15-30 years ( $N=4,180$ ), 30-45 years ( $N=10,062$ ), 45-60 years ( $N=18,805$ ), and 60 or more years ( $N=22,504$ ). Further, summary data from cohorts of European ancestry only were meta-analyzed and used in subsequent analyses. We discarded variants present in less than three studies. Approximately 23,700,000 markers (including SNPs and INDELS) were assessed for association. We applied the conventional genome-wide significance level (GWS,  $P < 5 \times 10^{-8}$ ) for SNP discovery.

### Assessment of Age-dependent effects

We selected SNPs which were suggestively (12,567 SNPs,  $P < 5 \times 10^{-6}$ ) associated with BMD in the overall meta-analysis, present in at least 2 studies per age-bin and with MAF differences across these meta-analyses lower than 0.5. We clumped this dataset with an  $r^2 \geq 0.8$ , using as reference the most strongly associated SNPs with BMD and, pruning remaining SNPs within 0.7 Mb of each other. Age-dependent effects were assessed using a meta-regression approach for 1,464 SNPs obtained after this selection procedure. We ran a linear regression of the SNP effect estimates onto an intercept and the median age of each subgroup (e.g., each study stratified in age-bins). As proposed previously<sup>20</sup>, standard errors of the effect estimates of each subgroup were multiplied by the square root of the genomic inflation factor when it was greater than 1.

We performed the meta-regression using the Metafor package<sup>21</sup>, and any statistical evidence of linear association was corrected for multiple testing (Bonferroni correction;  $0.05/1,464 = 3.4 \times 10^{-5}$ ). The difference between beta-estimates in children vs. elderly meta-analyses (Pdiff) was tested using Easy-strata<sup>22</sup>.

#### Approximate conditional meta-analyses

Conditional analyses were undertaken based on the meta-analysis of the studies of European ancestry only (N=56,284). Only variants in the loci that reached GWS in this meta-analysis were assessed. The Rotterdam Study I (n=6,291) was used as reference for precise calculation of the linkage disequilibrium (LD) between the analyzed markers. We used an iterative strategy as implemented in GCTA<sup>23</sup> to determine: 1) independence of association signals within loci discovered in our study, by means of stepwise model selection procedure per chromosome (--massoc-slct routine); and 2) the novelty of the association signals discovered by our meta-analysis with regard to variants reported in previous well-powered GWAS of different bone traits (**Table S3**). To this end, we performed the association analysis conditional on 78 variants present in our data and associated with different bone-traits (--massoc-cond routine). These 78 SNPs were selected from different GWAS publications<sup>6-10;12-14</sup>, assuring their independence to avoid collinearity issues.

## *Shared Genetic architecture of TB-BMD fracture and other traits*

### LD score regression analyses

We used the LD score regression package to estimate the heritability of TB-BMD and rule out that our results were a product of bias (e.g., residual population stratification or cryptic relatedness). LD score regression uses GWAS summary statistics and assesses the SNP-heritability based on the expected relationship between linkage disequilibrium (LD) of neighboring SNPs and strength of association under a polygenic model<sup>24</sup>. As this methodology relies on the LD structure throughout the genome, we restricted this analysis to summary statistics from the meta-analysis of cohorts comprising only individuals from European ancestry. We used the publicly available, pre-computed LD structure data files specific to European populations of the HAPMAP 3 reference panel. An extension of this method allows estimating the genetic correlation between two traits<sup>25</sup>. This can be performed in the LDhub pipeline, a web utility which gathers data from many different GWAS meta-analysis<sup>26</sup>. From the 199 traits, currently available in the website, we have restricted our analysis to those traits whose heritability z-scores were larger than 4 and were analyzed only in European ancestry individuals (following the recommendations in the LD score software website (**Web Resources**)). Additionally, we incorporated data from a recent GWAS meta-analysis of any-type of fracture in individuals from European ancestry (N= 264,267; 37,778 cases) (K.T, unpublished data). In total, we assessed the genetic correlation between TB-BMD and 74 traits.

## **Mendelian randomization analysis**

We undertook a two-sample Mendelian randomization approach<sup>27</sup> to estimate the causal effect of TB-BMD on any-type of fracture in the Europeans samples. In short, we constructed a score based on the independent genetic variants from the TB-BMD meta-analysis (European set and excluding secondary signals), whenever the selected variant was not present in the fracture meta-analysis, the second variant with the lowest p-value in the locus ( $P < 5 \times 10^{-8}$ ) and  $r^2 > 0.8$  was used as proxy. Thereafter, estimates derived from the TB-BMD summary statistics were pooled using methods similar to inverse-variance weighted fixed meta-analysis using the meta R-package (**Web Resources**).

## ***Search for biological and functional knowledge of the identified association regions***

For all those SNPs outside a 500Kb window from previously known bone associated SNPs we did a literature search in PubMed and Web of Science to evaluate if nearby genes (within 500Kb) were known to play a role in bone metabolism. Also, we determined if the annotated genes underlie any human Mendelian disorder with a skeletal manifestation, had knockout mouse models with a skeletal phenotype or were annotated to pathways critical to bone metabolism. Genomic annotation for all SNPs was made based on UCSC hg19.

## **DEPICT analyses**

We used DEPICT<sup>28</sup>, a recently developed tool to prioritize genes at the associated regions, define possible pathways by enrichment testing, and identify tissue and cell types in which genes from loci associated with TB-BMD. The methodology first selects all lead SNPs below a certain threshold with respect to a target P-value. We tested both the complete set of GWS

SNPs and the subset of those mapping only to loci not previously reported. Enriched gene-set were group based on the degree of gene overlap into 'meta gene-sets' as proposed earlier<sup>29</sup>, and their correlation visualized using Cytoscape 3.4 (**Web Resources**).

#### **Functional annotation to microRNA binding sites**

We used the PolymiRTS<sup>29</sup>, miRdSNP<sup>30</sup>, and microSNiPer<sup>31</sup> databases to obtain a list of variants located in predicted microRNA binding sites on the 3'UTRs of genes, as described in detail elsewhere<sup>32</sup>. In summary, index SNPs (most associated variant) of the GWS loci were submitted to SNAP (**Web Resources**) to retrieve their high LD proxy SNPs (with  $r^2 > 0.8$ , limit distance 500 kb, and CEU panel) in the 1000 genomes project. The resulting list of SNPs was annotated to the list of microRNA binding site variants obtained from the above mentioned publicly available databases.

#### **Functional enrichment analysis of trait-associated variants**

GWAS Analysis of Regulatory or Functional Information Enrichment with LD correction (GARFIELD)<sup>33</sup> was used to characterize the putative functional contribution of TB-BMD associated variants mapping to non-coding regions. GARFIELD employs a non-parametric analysis to calculate fold enrichment values for regulatory marks, at given significance thresholds and then tests them for significance via permutation testing while accounting for LD, MAF and local gene density<sup>33</sup>. We used data regarding DNase I hypersensitive sites, transcription factor binding sites, histone modifications and chromatin states (ENCODE and Roadmap Epigenomics) from 424 cell types and tissues to capture and characterize possible cell-type-specific patterns of enrichment, as provided in the GARFIELD software (**Web**

**Resources**). Fold enrichment statistics were tested at the four different significance thresholds (i.e.,  $1 \times 10^{-8}$ ,  $1 \times 10^{-7}$ ,  $1 \times 10^{-6}$  and  $1 \times 10^{-5}$ ). Multiple-testing correction was performed on the effective number of annotations used, using the default P-value threshold of  $1 \times 10^{-4}$ .

## **Knockout animal models and gene expression in bone cells**

### **Animal models survey**

We surveyed databases from *The International Mouse Phenotyping Consortium*<sup>34</sup> together with *The International Knockout Mouse Consortium*<sup>35</sup> to identify knockout models of candidate genes resulting in skeletal phenotypes. Furthermore we mined data from *The Origins of Bone and Cartilage Disease* (OBCD) project<sup>36</sup>, specialized in murine skeletal phenotypes including Digital X-ray microradiography on femurs and tail vertebrae, Micro-CT analysis, femur three-point bend test load–displacement curves and tail vertebrae compression testing from knockout mice and wild-type controls at 16 weeks of age.

### **Gene expression in murine bone cells**

Gene expression profiles of candidate genes were examined in primary mouse osteoblasts undergoing differentiation and bone marrow derived osteoclasts. To study murine osteoblasts, pre-osteoblast-like cells were obtained from neonatal calvaria collected from C57BL/6J. Next Generation RNA sequencing using an Illumina HiSeq 2000 was used to evaluate the transcriptome every two days from day 2 to 18 days post osteoblast differentiation<sup>7</sup>. Expression of genes in murine osteoclasts was determined using publicly available data obtained using Next-Gen RNA-sequencing applied to bone marrow derived osteoclasts obtained from 6-8 week old C57BL/6 mice<sup>37</sup>.

## Gene expression in human bone cells

Gene expression profiles of candidate genes were examined in human bone marrow derived mesenchymal stem cells differentiated into osteoblast. Total RNA (n=3) was isolated at day 0 (MSCs) and day 4 of osteoblast differentiation<sup>38</sup>. Also, RNA was isolated during osteoclast differentiation. Peripheral blood mononuclear cells derived from buffy coats (Sanquin, Amsterdam, the Netherlands) were seeded in 96-well plates ( $5 \times 10^5$  cells per well) as previously described<sup>39</sup>. Total RNA (n=3) was isolated using Trizol at day 0 (PBMCs) and at day 7 of osteoclast differentiation. Illumina HumanHT-12 v3 BeadChip human whole-genome expression arrays were used for expression profiling. The quality of isolated RNA was assessed on a 2100 Bioanalyzer (Agilent Technologies). Data were analyzed as described in detail previously<sup>38</sup>. Genes were designated as being expressed when at least one probe coding for the gene was significantly present in at least 2 of the 3 biological replicates.

## **Results**

### ***TB-BMD GWAS meta-analyses***

#### **Analyses including all age-strata**

Our meta-analysis of TB-BMD GWAS summary statistics (N=66,628) identified variants in 76 independent loci associated with TB-BMD at a genome-wide significant (GWS,  $P \leq 5 \times 10^{-8}$ ) level (**Figure 1, Table S4**). Overall, there was no evidence of a strong inflation (genomic inflation factor ( $\lambda$ ) of 1.08, **Figure S1**). Yet, inflation was observed in the range of common variants ( $0.2 > \text{MAF} < 0.5$ ,  $\lambda = 1.19$ ) due to polygenicity (LD score regression intercept = 1.007). In our results, one of the signals mapping to *LDLRAD3* was driven entirely by individuals of African

background (MAF=0.043 in YRI panel) since the two associated variants are monomorphic in all other populations. The low allele frequency of this variant in our study (MAF= 0.025) and our limited statistical power (N=6,748) in non-European samples warrants independent replication efforts to exclude the possibility of a false-positive association.

In addition, a meta-analysis comprising 56,284 individuals of European ancestry (~84% of the study population) identified variants in two additional GWS loci (**Figures S1-S2, Table S5**). Association signals mapping to these loci were close to the GWS threshold in the overall meta-analysis ( $P=1\times 10^{-7}$ ) and showed no evidence of heterogeneity ( $P_{\text{het}}>0.1$ ). One of them, in 12q24.21 (*MED13L*), has not been previously associated with bone parameters (**Table 1, Figure S3**), while the other in 21q22.13 (*CLDN14*), is not fully independent from the previously reported hip-BMD association signal<sup>13</sup> (**Table S5**).

Of the 78 identified loci, variants in 35 (45%) were not located within 500 kb of known association signals nor in regions of extended LD with them (**Table 1, Figure S4**). Index SNPs at these 35 loci were, in general, common non-coding variants. Twenty-two of these, are located in close proximity to genes likely to influence bone metabolism as shown by previous functional studies (**Table 1, Figure S3**), including *CSF1* ([MIM 120420] important for osteoclast differentiation<sup>40</sup>) and *SMAD3* ([MIM 603109] a critical component of the TGF-beta signaling pathway<sup>41</sup>). Across these 35 signals, 31 of the index SNPs were nominally associated ( $P<0.05$ ) with either lumbar spine or femoral neck BMD in the same direction as in the previously published GEFOS GWAS meta-analysis<sup>7</sup> (**Table 1**). This comparison was not possible for the rs113964474 variant, because it was not available in the GEFOS study. Moreover, we found directionally-concordant effect estimates ( $P < 0.05$ ) for 73 of the 78 index SNPs of known bone



association signals (**Table S3**). The markers which failed to replicate in our study were either previously associated with lumbar spine BMD but not femoral neck BMD (rs3905706 [*MPP7*, 10p12.1] and rs1878526 [*INSIG2*, 2q14.2]), associated specifically with the hip trochanter and intertrochanteric subregions (rs1949542 [*RP11-384F7.1*, 3q13.32]), or associated with BMD only in women (rs7017914 [*XKR9*, 8q13.3]) or only in children (rs754388 [*RIN3*, 14q32.12]).

### **Age-dependent effects**

Meta-analyses across age strata resulted in the identification of variants mapping to 2 additional loci that were not detected in the overall meta-analysis (**Figure S5; Table S6**). In children (age group 0-15 years), the previously known 14q32.12 locus<sup>8</sup>, harboring *RIN3* (rs72699866,  $P=1\times 10^{-8}$ ); and in the middle-aged (age group 45-60 years), a signal in the 19q12 locus mapping in the vicinity of *TSHZ3* (rs6510186,  $P=3.1\times 10^{-8}$ ) were identified. The rs72699866 variant leading the *RIN3* signal in the youngest age stratum showed no evidence of association ( $P=0.16$ ) and high heterogeneity ( $P_{\text{het}}=6.6\times 10^{-5}$ ) in the overall meta-analysis. In fact, the effect of rs72699866 decreased significantly with age ( $P_{\text{trend}}=1.69\times 10^{-9}$ ) (**Figure S6**) and showed a significant difference between the two extreme groups, i.e. children vs elderly ( $\beta_{0-15}=0.099$  [0.066, 0.134];  $\beta_{>60}=-0.035$  [-0.060, -0.010];  $P_{\text{diff}}=4.32\times 10^{-10}$ ). In contrast, the rs6510186 variant [19q12] showed nominal evidence of association and heterogeneity in the overall meta-analysis ( $P=0.02$ ;  $P_{\text{het}}=0.03$ ). Nevertheless, no clear pattern of age-dependency was identified ( $P=0.2$ ) for this SNP (**Figure S6**).

We also applied meta-regression analysis and found that variants mapping to 42 different loci showed nominally significant age dependent effect ( $P<0.05$ ) (**Table S7, Figure S7**). In summary,

27 (64%) of the loci showed stronger effects in the older age groups. Of these, variants in the 6q25.1 (*ESR1*) and 13q14.11 (*RANKL*) loci remained significant after multiple-testing correction ( $P < 3.4 \times 10^{-5}$ ) (**Figure 2**); while variants in 6p21.1 (*RUNX2*, rs148460475), 15q21.2 (*CYP19A1*, rs2414098), 17q21.31 (*MEOX1*, rs74835612) and 11p15.1 (*SOX6*, rs11822790) were only suggestive at  $P < 1 \times 10^{-3}$ .

#### **Conditional association analyses**

The step-wise conditional approach included studies comprising only individuals of European ancestry, as the method used relies on appropriate representability of the LD reference. Of the 76 GWS loci identified in the overall analysis, variants in 57 (19 previously unreported) loci were also GWS in the European-only analysis (**Figure S2**), likely a consequence of the lower power in this subgroup. We identified 81 SNPs independently associated with TB-BMD mapping to 58 different loci (one European-specific), 18 of which depicted multiple distinct signals attaining GWS (**Table S8**). These independent variants together explained 10.2% of TB-BMD variance. This proportion is slightly higher than the 7.4% TB-BMD variance explained by the 78 known variants associated with bone traits. Moreover, we identified independent signals in 13 of the 78 known bone loci after conditional analyses. (**Figure S2; Table S8**).

#### ***Shared Genetic architecture of TB-BMD, fracture and other traits***

SNP-heritability of TB-BMD in the European samples was estimated to be 0.259 (SE 0.017). TB-BMD was highly genetically correlated with BMD measured at other skeletal sites ( $\rho > 0.9$ ). Among the non-BMD traits, all-type of fracture showed the highest correlation [ $\rho = -0.61$  ( $P = 1.6 \times 10^{-27}$ )]. The MR approach indicated a strong causal relation where per 1 standard

deviation decrease in genetically determined TB-BMD there is 56% increase in the risk of fracture (Odds ratio 1.56 [1.50-1.62]). Other anthropometric, metabolic and disease traits showed significant (yet weak) correlation with TB-BMD (**Table S9, Figure 3**). In contrast, other established risk factors for osteoporosis such as menopause or age of menarche showed no significant genetic correlation with TB-BMD.

#### ***Biological and functional knowledge of the genes in BMD-associated loci***

Loci not previously reported and their potential role in bone metabolism are summarized in **Table 1**. Several loci harbor genes implicated directly in bone metabolism (*SLC8A1* [MIM 182305], *PLCL1* [MIM 600597], *ADAMTS5* [MIM 605007]), affecting osteoblast or osteoclast differentiation and activity (*CSF1* [MIM 120420], *DUSP5* [MIM 603069], *SMAD3* [MIM 603109], *SMAD9* [MIM 603295], *CD44* [MIM 107269]), participating in Wnt signaling (*FZD7* [MIM 603410], *TCF7L1* [MIM 604652]), or regulating processes such as manganese or calcium absorption (*GCKR* [MIM 600842], *DGKD* [MIM 601826], *SLC30A10* [MIM 611146]) among others<sup>40-61</sup>; while genes in at least 14 loci exert a potential novel role in bone biology. Rodent knockout models of several genes in the implicated loci, show an altered skeletal phenotype (e.g., osteopetrosis [*Csf1*<sup>40</sup>], increased bone resorption [*Aqp1*<sup>50</sup>, *Cyp19a1*<sup>57</sup>, *Cd44*<sup>53</sup>], impaired skeletogenesis [*Apc*<sup>49</sup>, *Runx1*<sup>60</sup>, *Smad3*<sup>41</sup>], deformities in the axial skeleton [*Btg1*<sup>62</sup>, *Atpaf2*<sup>63</sup>]). Whereas an effect on bone can be inferred for genes in other associated loci, for example, *CYP19A1* [MIM 107910] in 15q21.2 is an estrogen synthesis gene, being estrogen a key compound for bone maturation and maintenance, and *ZKSCAN5* [MIM 611272] in 7q22.1 is associated with circulating dehydroepiandrosterone sulphate (DHEAS) levels<sup>51</sup>. DHEAS levels are positively correlated with BMD in adults and post-menopausal women<sup>64</sup>. Across these loci,

not previously reported as associated with BMD variation, we identified six exonic variants associated with TB-BMD, three of which were nonsynonymous variants all cataloged as benign both by SIFT and polyphen2. We also identified 53 GWS coding variants in known loci, of which 33 are non-synonymous (**Table S10**). Only a low-frequency variant in *LRP5* [MIM 603506], rs4988321/A (11:68174189, MAF=0.04), has a clinical annotation, constituting a homozygous G-to-A transition variant identified in a person with osteoporosis-pseudoglioma syndrome (OPPG [MIM 259770])<sup>65</sup>.

#### **DEPICT analyses**

Based on the overall meta-analysis, 53 genes were prioritized (FDR<0.05), 15 of them mapping to loci not previously described (**Table S11**). Cells and tissues from the musculoskeletal system presented the largest enrichment of gene expression within the associated loci (**Figure 4**). These genes were overrepresented in 182 pathways clustered in 25 ‘meta gene-sets’ (**Table S12**). The large majority of the clusters are involved in musculoskeletal development and bone homeostasis (**Figure 4**). The most significant of these implicated the regulation of cell growth, and the TGFB signaling pathway and its mediating SMAD proteins.

Restricting the DEPICT analysis to the subset of not previously reported associated regions resulted in significant enrichment of genes expressed in the musculoskeletal and immunological systems (**Figure S8**). Genes mapping to these loci were overrepresented in the SMAD binding pathway and TGFBR2 PPI (protein-protein interaction) subnetwork (FDR<0.05).

#### **Functional annotation to microRNA binding sites**

We then assessed if the index SNPs of the 80 GWS loci detected in the main and subsequent GWAS (or their proxies in strong LD;  $r^2 > 0.8$ ) were located in predicted microRNA binding sites within the genes' 3'UTRs and thus, were expected to disrupt the regulation of gene expression (**Table S13**). The index SNP within the 3'UTR of *ZKSCAN5* (mapping to a locus not previously identified), rs34670419 (MAF=0.04), is predicted to create a binding site for miR-382-3p, a microRNA which is expressed in osteocytes and has been recently shown to be involved in osteogenic differentiation<sup>66</sup>. In addition, eight proxy SNPs (mapping to *PSMD13*, *ABCF2*, *GALNT3*, *PKDCC*, *REEP5*, *PPP6R3*, *AAGAB* and *TOM1L2*) are predicted to influence the binding of microRNAs to transcripts of their host gene.

#### **Functional enrichment analysis of trait-associated variants**

As typically found in GWAS, the great majority of identified associations emerged from non-coding common variants and hold no direct annotation to molecular mechanisms. To assess if there is relative enrichment of regulatory genomic marks underlying the associated variants in a cell-specific context, we used GARFIELD<sup>33</sup>. We found relative ubiquitous enrichment for TB-BMD variants (Empirical  $P < 2.4 \times 10^{-4}$ ) in DNase I hypersensitive sites across the different cell types (**Figure S9**). Further, we found higher levels of fold-enrichment for enhancers (median 3.6, range [2.7, 4.4]) and promoters (median 3.2, range [2.9, 3.5]) than for transcribed regions (median 1.8, range [1.5, 2.2]).

#### **Gene expression in bone cells and knockout animal models**

From the 53 genes prioritized by DEPICT only 49 had a mouse orthologue (**Table S14**). From these genes, only *Mepe* (osteocyte-specific) and *Foxl1* were not expressed in murine osteoblast

or osteoclast. Moreover, 61% of the prioritized genes were expressed in human cells *in vitro* during osteoblast or osteoclast differentiation (**Table S14**). *AQP1* was the only prioritized gene mapping to a locus not previously reported showing no expression in the human bone cells differentiation experiments.

Knockout models were widely available in at least one of the different databases assessed. Nevertheless in-depth bone phenotyping performed under the OBCD project was only available for four knockout models (**Table S15**). Two of these, *DUSP5* and *CD300LG* showed no significant bone phenotype. The *TCF7L1* knockout model only showed lower cortical diameter in the femur without other clear bone phenotype. Nevertheless, *TCF7L1* was shown to be expressed during osteoblastogenesis. Conversely, homozygous knockout for *CREB3L1* showed a clear bone phenotype consisting of low BMC both at the vertebrae and femur together with a strong trabecular and cortical phenotype affecting bone strength (**Figure S10**). *CREB3L1* maps to 11p11.2, a previously identified BMD locus<sup>6</sup> harboring *ARHGAP1* and *LRP4* as candidates to underlie the GWAS signal in a region of extended LD.

## Discussion

This meta-analysis of TB-BMD comprising up to 66,000 individuals identified variants in 36 loci not previously reported and replicated at GWS level several association signals identified by GWAS of diverse bone phenotypes. Bioinformatics analyses suggest enrichment of these 36 loci for genes expressed in the musculoskeletal system, and solidly represented in the SMAD binding pathway and the TGFBR2 PPI subnetwork. We also demonstrate that for variants in few loci the size of the effect is age dependent; variants in two loci (*RIN3* and *TSHZ3*) were

identified only by the age-stratified analyses despite less power (smaller sample size); while for variants in two other loci (*ESR1* and *RANKL*) there was significant evidence of age heterogeneity derived from a meta-regression of the genetic effects with age. Our results strengthen the evidence that genetic variants influence BMD from a young age and support the value of peak bone mass as an important determinant of bone health later in life.

Traditionally, DXA-BMD measurements performed at sites of high fracture risk (i.e., femoral neck, lumbar spine and forearm) have been used in genetic epidemiological investigations of bone health in adults. Instead, we have used BMD measurements derived from total body scans. Not only do we show a high overlap of association signals with previous GWAS of different bone traits, including DXA, pQCT and ultrasound measurements, but we have also identified unreported loci. Five known associations failed to replicate in our studies, even though we cannot discard these associations constitute false-positives, these results might also indicate that variants whose effect is highly specific to skeletal sites, skeletal properties, sex or age groups cannot be detected in our TB-BMD meta-analysis. It is plausible that more variants of this type exist and will be discovered as site-specific BMD meta-analyses are performed in increasingly powered settings. Furthermore, the genetic correlation of TB-BMD with BMD measured at other sites was close to one. Whilst, we found that a decrease of one standard deviation in the genetically determined TB-BMD resulted in at least 50% higher odds of suffering a fracture. Significant genetic correlations with other traits (i.e., BMI, IGF1 and ulcerative colitis) reflect the systemic context of skeletal biology and merit further study by future efforts to elucidate the underlying mechanisms.

505 Genes in the associated loci were highly expressed in the musculoskeletal system and  
506 overrepresented in gene-sets related to bone development. The prioritized gene *CREB3L1* [MIM  
507 616215] in 11p11.2 observed a clear bone phenotype in our mouse knockout model, which  
508 corroborates the findings of previous work showing substantial rescue of *CREB3L1* deficiency  
509 with bisphosphonates and its critical role for bone formation<sup>67</sup>. This locus characterized by  
510 extended LD, also harbors *LRP4* [MIM 604270] whose knockout model presents with increased  
511 trabecular and cortical bone mass<sup>68</sup>. This is in line with our conditional analysis identifying  
512 multiple independent signals in the region making it likely that both genes are influencing bone  
513 biology. Altogether, we demonstrated that TB-BMD offers a powerful alternative to identify  
514 genetic variants associated with bone metabolism.

515 Variants mapping to 14q32 harboring *RIN3* [MIM 610223] were only associated at a GWS level  
516 in children (i.e., <15 years), and were only nominally significant in the elderly group (i.e., >60  
517 years). This age-related heterogeneity may explain why this locus has not been detected in  
518 BMD meta-analyses in adults, although being identified in relation to pediatric BMD<sup>8</sup> and  
519 Paget's disease (PDB [602080]) GWAS<sup>69</sup>. In addition, another signal mapping to 19q12  
520 harboring *TSHZ3* [MIM 614119] was significant in adults aged 45-60 years but not in other age  
521 groups analyzed or in previous studies, alluding to a false-positive association, thus replication  
522 of this finding is necessary.

523 Our analyses revealed variants in the 6q25.1 (*ESR1*) and 13q14.11 (*RANKL*) loci demonstrating  
524 the most compelling evidence for age-modulation effects. The 6q25.1 locus harboring *ESR1*  
525 [MIM 133430], an important genetic factor in normal BMD variability, was not associated with  
526 BMD in children below 15 years of age, where the largest cohorts (i.e., Avon Longitudinal Study



527 of Parents and Children (ALSPAC) and the Generation R Study) comprise predominantly pre-  
528 pubertal children. As levels of estradiol before puberty are low<sup>70</sup>, a negligible effect of *ESR1*  
529 variants on BMD is expected. Likewise, in mouse models the expression of *RANKL* [MIM  
530 602642] in bone is markedly increased with advancing age from young to adult and related to  
531 bone loss<sup>71</sup>. Accordingly, variants influencing *RANKL* expression show a larger effect later in life.  
532 In general, a substantial heterogeneity of the genetic effects in the overall meta-analysis was  
533 explained by age, nevertheless, the inclusion of larger sample sizes (avoiding age exclusion  
534 criteria and incrementing statistical power) leveled off the loss of power due to the  
535 heterogeneity of the genetic effects.

536 In brief, variants with evidence of age-specific effects were exceptional in our study. These  
537 results might reflect a lack of statistical power as only SNPs showing suggestive evidence  
538 ( $P < 5 \times 10^{-6}$ ) of association with TB-BMD in the overall meta-analysis were tested for age-specific  
539 effects. This selection criteria aimed to include SNPs whose heterogeneity might have  
540 hampered their statistical significance in the overall meta-analysis, and at the same time  
541 maximize the power to discover variants with real age-dependent effects. Alternatively, these  
542 results indicate that most of the genetic variants identified so far, by us and others, influence  
543 BMD from early ages onwards, and their effect persist throughout the life course. However,  
544 variants in 27 of the 42 loci (64%) showing nominal evidence for age dependent effects had  
545 larger effects in the older groups. Nonetheless, this requires careful interpretation given the  
546 uneven sample sizes between the age groups and the criteria to select markers for the meta-  
547 regression based on significance in the overall meta-analysis. Collectively, this argues in favor of

enlarging studies focused on younger populations –where the statistical power is still restricted – to discover additional genetic variants influencing BMD.

Our study has some limitations. A key disadvantage of our design is that we group the data based on age spans rather than life stages. Crucial information for this assesment, such as puberty onset in children and adolescents or menopausal status in the adults, was not available across the majority of the cohorts. Other strategies like using shorter age spans will resulted in even less statistical power of the discovery setting. Similarly, despite the large sample size of our study, we identified very few variants in the low-frequency spectrum (MAF <5%) indicating that comprehensive surveys of rare variation influencing BMD still require even larger sample sizes, on top of better resources for imputation of the rarer variants, possibly needing population-specific references. Such strategies will be key to explain a larger fraction of the genetic variability of BMD phenotypes, as illustrated for other traits such as height or BMI<sup>72</sup>. Moreover, the identified SNPs are in their vast majority, non-coding variants, raising the possibility that the causal genes are different from the candidate genes we have prioritized based on the current biological knowledge and bioinformatic prediction tools. Additional functional studies are required to determine the potential role of the genes in the identified loci.

In conclusion, we performed a genome-wide survey for association with DXA derived TB-BMD, combining data from five age groups including children and older individuals. In contrast to previous large-scale meta-analyses<sup>6;7</sup>, we used DXA derived TB-BMD rather than measurements on specific skeletal sites prone to fracture to identify genetic factors influencing BMD variation. We demonstrate that TB-BMD is a valid phenotype for this purpose, as we replicated more than

90% of the previously reported signals. Most importantly, we identify variants in 36 loci associated with TB-BMD not previously reported by previous GWAS of bone phenotypes. Our results show steadiness in the magnitude of the genetic effects on BMD for most of the BMD-associated variants. While the contrasting skeletal physiology across different age periods is well established (i.e. endochondral ossification, linear growth, modelling, remodeling, etc.), peak bone mass acquisition remains the major determinant of variability at any age. These findings strongly support the importance of the bone accrual process in the definition of BMD status and fracture susceptibility throughout the life course.

#### **Accession Numbers**

GWAS Summary data for the main and age-strata meta-analyses together with the corresponding regional plots of GWS signals have been deposited in the GEFOS website (**Web Resources**). Gene expression data presented in this paper can be retrieved from the Gene Expression Omnibus (GEO) as follows: Murine osteoclasts (GSM1873361) and osteoblasts (GSE54461); human osteoblast differentiation (GSE54461).

#### **Supplemental Data**

Supplemental data include a full list of acknowledgements, cohort short descriptions, 15 tables and 10 figures.

#### **Acknowledgements**

The authors would like to thank the many colleagues who contributed to collection and phenotypic characterization of the clinical samples, as well as genotyping and analysis of the

590 GWAS data. Part of this work was conducted using the UK Biobank resource.

591 **Conflict of interests**

592 Psaty serves on the DSMB of a clinical trial for the manufacturer (Zoll LifeCor) and on the

593 Steering Committee of the Yale Open Data Access Project funded by Johnson & Johnson.

594

595 **Web Resources**

596 GARFIELD, <http://www.ebi.ac.uk/birney-srv/GARFIELD/>GEFOS, <http://www.gefos.org/>

597 LDhub, <http://ldsc.broadinstitute.org/>

598 Meta R-package, <https://github.com/guido-s/meta>

599 OBCD, <http://www.boneandcartilage.com/>

600 OMIM, <http://www.omim.org/>

601 SNAP, <http://archive.broadinstitute.org/mpg/snap/>

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## Figure Titles and Legends

**Figure 1. Manhattan plot of association statistics ( $-\log_{10}(P\text{-values})$ ) for TB-BMD overall meta-analysis.** Each dot represents a SNP and the x-axis indicates its chromosomal position (built 37 NCBI). Red dots represent SNPs at GWS loci that are not within  $\pm 500\text{Kb}$  of leading SNPs in previous GWAS with different bone traits. Dashed horizontal red and yellow lines mark the GWS threshold ( $P < 5 \times 10^{-8}$ ) and suggestive threshold ( $P < 1 \times 10^{-6}$ ), respectively. Novel loci in the only-CEU analysis are not shown.

**Figure 2. Age dependence of the genetic variant effect in the meta-regression.** The panels display leading SNPs from two loci exhibiting significant evidence for age influences. Heterogeneity P-values ( $P_{\text{het}}$ ) are reported for the overall meta-analysis. In the left panels, each circle represents a study subgroup (i.e., study divided in age strata), with the circle size proportional to the inverse variance of the SNP main effect. In the right panels, forest plots display estimates obtained from each age-bin meta-analysis, with the symbol size proportional to the inverse variance of the SNP main effect.

**Figure 3. Genetic correlations between TB-BMD and other traits and diseases.** Calculation was based on the summary statistics of the only-European meta-analysis ( $N=56,284$ ) and estimated by LD score regression implemented in LDHub. The diagram only show traits whose correlation with TB-BMD was significant ( $P < 0.05$ ).

**Figure 4. Depict results for gene-set and cell/tissue enrichment analyses.** **Top panel:** 25 Meta gene-sets were defined from similarity clustering of significantly enriched gene sets ( $\text{FDR} < 5\%$ ). Each Meta gene-set was named after one of its member gene sets. The color of the Meta gene-sets represents the P-value of the member set. Interconnection line width represents the Pearson correlation  $\rho$  between the gene membership scores for each Meta gene-set ( $\rho < 0.3$ , no line;  $0.3 \leq \rho < 0.5$ , narrow width;  $0.5 \leq \rho < 0.7$ , medium width;  $\rho \geq 0.7$ , thick width). **Bottom panel:** Bars represent the level of evidence for genes in the associated loci to be expressed in any of the 209 Medical Subject Heading (MeSH) tissue and cell type annotations. Highlighted in orange are these cell/tissue types significantly ( $\text{FDR} < 5\%$ ) enriched for the expression of the genes in the associated loci.

## Tables

**Table 1. Index SNPs of loci not previously associated with BMD.** Variants associated with TB-BMD in the all-ages combined meta-analysis that map outside +/- 500 Kb of known index SNPs of genetic associations with different bone traits. Genomic coordinates are on build 37 of the human genome. Notes refer to annotation based on the closest gene. Associations with Lumbar Spine (LS) and Femoral Neck (FN)-BMD <sup>10</sup>. Beta coefficients and allele frequencies (EAF) are reported for the A1 allele

| CHR | BP        | rsnumber     | Locus    | A1 | A2 | EAF  | Effect | P        | N     | annotation | closest gene    | Notes   | LS-beta | LS-P     | FN-beta | FN-P     |
|-----|-----------|--------------|----------|----|----|------|--------|----------|-------|------------|-----------------|---|---------|----------|---------|----------|
| 1   | 8422676   | rs2252865    | 1p36.23  | T  | C  | 0.32 | -0.033 | 4.72E-08 | 66075 | intronic   | <i>RERE</i>     | Novel biology   | -0.019  | 0.043    | -0.025  | 0.002    |
| 1   | 110475971 | rs7548588    | 1p13.3   | T  | C  | 0.61 | -0.037 | 9.29E-09 | 66240 | intergenic | <i>CSF1</i>     | Osteoclast differentiation <sup>40</sup>                        | -0.030  | 0.001    | -0.022  | 0.005    |
| 1   | 220038825 | rs185048405  | 1q41     | T  | C  | 0.54 | 0.042  | 3.07E-09 | 66540 | intronic   | <i>SLC30A10</i> | Manganese transport <sup>42</sup>                               | -0.035  | 0.076    | -0.003  | 0.878    |
| 2   | 27741072  | rs780096     | 2p23.3   | C  | G  | 0.44 | -0.031 | 4.58E-08 | 66578 | intronic   | <i>GCKR</i>     | Calcium regulation <sup>43</sup> , hepatic traits <sup>44</sup> | -0.014  | 0.129    | -0.017  | 0.029    |
| 2   | 40630678  | rs10490046   | 2p22.1   | A  | C  | 0.76 | 0.043  | 1.43E-10 | 65961 | intronic   | <i>SLC8A1</i>   | Bone mineralization <sup>45</sup>                               | 0.015   | 0.162    | 0.021   | 0.025    |
| 2   | 68962137  | rs10048745   | 2p13.3   | A  | G  | 0.25 | -0.039 | 6.44E-09 | 66565 | 5'-UTR     | <i>ARHGAP25</i> | Novel biology   | -0.050  | 1.03E-06 | -0.036  | 5.21E-05 |
| 2   | 85484818  | rs11904127   | 2p11.2   | A  | G  | 0.55 | -0.032 | 2.65E-08 | 66561 | intronic   | <i>TCF7L1</i>   | Factors in Wnt signaling <sup>46</sup>                          | -0.021  | 0.023    | -0.015  | 0.054    |
| 2   | 198874006 | rs1595824    | 2q33.1   | T  | C  | 0.47 | 0.034  | 2.65E-08 | 60171 | intronic   | <i>PLCL1</i>    | Negative regulation of bone formation <sup>47</sup>             | 0.022   | 0.201    | 0.052   | 2.20E-04 |
| 2   | 202799604 | rs2350085    | 2q33.2   | T  | C  | 0.87 | -0.064 | 3.80E-14 | 66412 | intergenic | <i>FZD7</i>     | Factors in Wnt signaling <sup>48</sup>                          | -0.042  | 0.002    | -0.044  | 1.96E-04 |
| 2   | 234303405 | rs838721     | 2q37.1   | A  | G  | 0.44 | -0.031 | 4.48E-09 | 65516 | intronic   | <i>DGKD</i>     | Calcium regulation <sup>43</sup>                                | -0.016  | 0.070    | -0.014  | 0.068    |
| 5   | 112221869 | rs818427     | 5q22.2   | T  | C  | 0.31 | 0.034  | 2.37E-08 | 66592 | intronic   | <i>APC</i>      | Bone metabolism <sup>49</sup>                                   | 0.004   | 0.645    | 0.008   | 0.327    |
| 5   | 122847622 | rs11745493   | 5q23.2   | A  | G  | 0.75 | 0.044  | 7.75E-12 | 66597 | promoter   | <i>CSNK1G3</i>  | Novel Biology   | 0.010   | 0.326    | 0.025   | 0.005    |
| 7   | 27989403  | rs757138     | 7p15.1   | T  | G  | 0.69 | -0.035 | 3.33E-08 | 66043 | intronic   | <i>JAZF1</i>    | Novel Biology   | -0.016  | 0.126    | -0.025  | 0.004    |
| 7   | 30957702  | rs28362721   | 7p14.3   | T  | C  | 0.18 | -0.059 | 6.71E-14 | 66274 | intronic   | <i>AQP1</i>     | Bone metabolism <sup>50</sup>                                   | -0.037  | 0.002    | -0.049  | 1.39E-06 |
| 7   | 50901491  | rs1548607    | 7p12.1   | A  | G  | 0.69 | 0.036  | 4.18E-08 | 66564 | intergenic | <i>GRB10</i>    | Novel biology   | 0.034   | 5.59E-04 | 0.005   | 0.517    |
| 7   | 99130834  | rs34670419   | 7q22.1   | T  | G  | 0.04 | -0.088 | 1.09E-08 | 66336 | 3'-UTR     | <i>ZKSCAN5</i>  | DHEAS and aging mechanisms <sup>51</sup>                        | -0.127  | 9.28E-08 | -0.080  | 8.19E-05 |
| 10  | 112245400 | rs73349318   | 10q25.2  | A  | T  | 0.87 | -0.047 | 2.68E-08 | 66341 | intronic   | <i>DUSP5</i>    | Osteoclast differentiation <sup>52</sup>                        | -0.042  | 0.001    | -0.051  | 8.76E-06 |
| 10  | 124015986 | rs10788264   | 10q26.13 | A  | G  | 0.48 | -0.034 | 2.61E-09 | 66565 | intergenic | <i>TACC2</i>    | Novel Biology   | -0.030  | 9.64E-04 | -0.029  | 1.29E-04 |
| 11  | 242859    | rs55781332   | 11p15.5  | A  | G  | 0.78 | -0.055 | 8.07E-16 | 66198 | intronic   | <i>PSMD13</i>   | Novel Biology   | -0.046  | 1.76E-05 | -0.026  | 0.005    |
| 11  | 35083633  | rs2553773    | 11p13    | C  | G  | 0.41 | -0.037 | 1.49E-10 | 66619 | intergenic | <i>CD44</i>     | Osteoclast activity <sup>53</sup>                               | -0.015  | 0.101    | -0.015  | 0.054    |
| 11  | 35981346  | rs113964474* | 11p.13*  | A  | G  | 0.03 | 0.485  | 1.41E-08 | 6748  | intronic   | <i>LDLRAD3</i>  | Novel Biology   | .       | .        | .       | .        |
| 11  | 69299537  | rs4980659    | 11q13.3  | C  | G  | 0.52 | 0.033  | 1.16E-08 | 66537 | intergenic | <i>CCND1</i>    | Target of Wnt signalling <sup>54</sup>                          | 0.039   | 1.58E-05 | 0.023   | 0.003    |
| 11  | 121913230 | rs725670     | 11q24.1  | A  | G  | 0.38 | -0.032 | 3.61E-08 | 66565 | intergenic | <i>BLID</i>     | Novel Biology   | -0.020  | 0.028    | -0.011  | 0.172    |
| 12  | 90334829  | rs10777212   | 12q21.33 | T  | G  | 0.35 | 0.045  | 5.05E-14 | 66619 | intergenic | <i>ATP2B1</i>   | Calcium absorption <sup>55</sup>                                | 0.028   | 0.003    | 0.021   | 0.010    |
| 12  | 116555786 | rs73200209** | 12q24.21 | A  | T  | 0.80 | 0.045  | 2.51E-08 | 51240 | intronic   | <i>MED13L</i>   | Novel biology   | 0.030   | 0.167    | 0.036   | 0.044    |
| 13  | 37487021  | rs556429     | 13q13.3  | A  | C  | 0.23 | 0.039  | 1.46E-08 | 66504 | intronic   | <i>SMAD9</i>    | Osteoblast differentiation <sup>56</sup>                        | 0.023   | 0.027    | 0.013   | 0.135    |
| 15  | 38340874  | rs12442242   | 15q14    | A  | G  | 0.85 | -0.051 | 4.94E-10 | 66403 | intergenic | <i>TMCO5A</i>   | Novel Biology   | -0.046  | 3.03E-04 | -0.047  | 2.26E-05 |
| 15  | 51537806  | rs2414098    | 15q21.2  | T  | C  | 0.39 | -0.033 | 1.99E-08 | 66562 | intronic   | <i>CYP19A1</i>  | Estrogen biosynthesis <sup>57</sup>                             | -0.034  | 0.007    | -0.038  | 0.001    |
| 15  | 67420680  | rs1545161    | 15q22.33 | A  | G  | 0.56 | 0.041  | 1.06E-12 | 66004 | intronic   | <i>SMAD3</i>    | Osteoblast differentiation <sup>41</sup>                        | 0.034   | 1.27E-04 | 0.035   | 5.78E-06 |
| 17  | 17804725  | rs8070128    | 17p11.2  | T  | C  | 0.58 | -0.039 | 1.98E-11 | 66625 | intronic   | <i>TOM1L2</i>   | Novel biology   | -0.033  | 4.80E-04 | -0.015  | 0.052    |
| 17  | 63771079  | rs9972944    | 17q24.1  | A  | G  | 0.41 | 0.036  | 6.87E-10 | 66595 | intronic   | <i>CEP112</i>   | Novel Biology   | 0.028   | 0.003    | 0.004   | 0.576    |
| 19  | 31654615  | rs6510186*** | 19q12    | T  | C  | 0.26 | 0.068  | 3.11E-08 | 18782 | intergenic | <i>TSHZ3</i>    | Novel Biology   | 0.004   | 0.713    | 0.006   | 0.492    |

|    |          |            |          |   |   |      |        |          |       |            |                |  |        |       |        |       |
|----|----------|------------|----------|---|---|------|--------|----------|-------|------------|----------------|--|--------|-------|--------|-------|
| 20 | 39103882 | rs6029130  | 20q12    | T | C | 0.30 | 0.035  | 3.50E-08 | 66497 | intergenic | <i>MAFB</i>    | Osteoclast differentiation <sup>58</sup> | 0.027  | 0.007 | 0.015  | 0.083 |
| 21 | 28773868 | rs1452102  | 21q21.3  | T | G | 0.59 | -0.035 | 1.74E-09 | 66489 | intergenic | <i>ADAMTS5</i> | Endochondral Ossification <sup>59</sup>  | -0.029 | 0.001 | -0.015 | 0.056 |
| 21 | 36970350 | rs9976876  | 21q22.12 | T | G | 0.45 | -0.038 | 8.01E-11 | 66514 | intronic   | <i>RUNX1</i>   | Osteoclast differentiation <sup>60</sup> | -0.019 | 0.031 | -0.016 | 0.041 |
| 21 | 40350744 | rs11910328 | 21q22.2  | A | G | 0.84 | -0.043 | 2.99E-08 | 66298 | intergenic | <i>ETS2</i>    | Osteoblast maturation <sup>61</sup>      | -0.028 | 0.020 | -0.028 | 0.007 |

\* Monomorphic in European cohorts. \*\* Reported statistics from the in the meta-analysis of European populations. \*\*\* Reported statistics from the meta-analysis in the 30-45 age-strata.